

Formulario del corso di Statistica I
Ingegneria Gestionale
a.a. 2009/10

- $\bar{x} = \frac{x_1 + \dots + x_n}{n}$, $S^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2$, $\widehat{Cov} = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})$,
 $r = \frac{\widehat{Cov}}{S_{x_i X} S_Y} = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sqrt{\sum_{i=1}^n (x_i - \bar{x})^2 \sum_{i=1}^n (y_i - \bar{y})^2}}$. $\sum_{i=1}^n (x_i - \bar{x})^2 = (\sum_{i=1}^n x_i^2) - n\bar{x}^2$, $\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y}) = (\sum_{i=1}^n x_i y_i) - n\bar{x}\bar{y}$.
- $n! = n \cdot (n-1) \cdot \dots \cdot 2 \cdot 1$, $0! = 1$. $\binom{n}{k} = \frac{n!}{k!(n-k)!} = \frac{n \cdot (n-1) \cdot \dots \cdot (n-k+1)}{k!}$.
- $P(A|B) = \frac{P(A \cap B)}{P(B)}$, $P(A \cap B) = P(A|B)P(B)$. A, B indipendenti:
 $P(A \cap B) = P(A)P(B)$, $P(A|B) = P(A)$, $P(B|A) = P(B)$. $P(A) = \sum_k P(A|B_k)P(B_k)$. $P(B|A) = \frac{P(A|B)P(B)}{P(A)}$.
- X discreta, valori a_j , $P(X = a_j) = p_j$, allora $E[X] = \sum_j a_j p_j$, $E[g(X)] = \sum_j g(a_j) p_j$, $E[X^2] = \sum_j a_j^2 p_j$. $P(X \in A) = \sum_{i: a_i \in A} P(X = a_i) = \sum_{i: a_i \in A} p_i$. $X \in \mathbb{N}$, $P(X \leq n) = \sum_{i=0}^n p_i$, $P(X \geq n) = \sum_{i=n}^{\infty} p_i$.
- X continua, densità $f(x)$, allora $E[X] = \int_{-\infty}^{\infty} x f(x) dx$, $E[g(X)] = \int_{-\infty}^{\infty} g(x) f(x) dx$, in particolare $E[X^2] = \int_{-\infty}^{\infty} x^2 f(x) dx$. $P(X \in A) = \int_A f(x) dx$.
- $Var[X] = \sigma_X^2 := E[(X - \mu_X)^2]$ dove $\mu_X = E[X]$. $Var[X] = E[X^2] - \mu_X^2$. $Cov(X, Y) = E[(X - \mu_X)(Y - \mu_Y)]$, $Cov(X, Y) = E[XY] - \mu_X \mu_Y$.
 $\rho(X, Y) = \frac{Cov(X, Y)}{\sigma_X \sigma_Y}$. $-1 \leq \rho(X, Y) \leq 1$.
- $E[aX + bY + c] = aE[X] + bE[Y] + c$. $Var[X + Y] = Var[X] + Var[Y] + 2Cov(X, Y)$. $Var[aX] = a^2 Var[X]$. Standardizzazione di X : $\frac{X - \mu_X}{\sigma_X}$.
- X, Y indipendenti: $P(X \in A, Y \in B) = P(X \in A)P(Y \in B)$. Implica $E[XY] = E[X]E[Y]$, $Cov(X, Y) = 0$, $\rho(X, Y) = 0$, $Var[X + Y] = Var[X] + Var[Y]$.
- $F(x) = P(X \leq x)$. $F(t) = \int_{-\infty}^t f(x) dx$. $F'(t) = f(t)$. $F(q_\alpha) = \alpha$.
- $\varphi(t) = E[e^{tX}]$, $\varphi'(0) = E[X]$, $\varphi''(0) = E[X^2]$; $\varphi_{aX}(t) = E[e^{taX}] = \varphi_X(at)$. X, Y indipendenti implica $\varphi_{X+Y}(t) = \varphi_X(t)\varphi_Y(t)$.
- $X \sim B(n, p)$: $P(X = k) = \binom{n}{k} p^k (1-p)^{n-k}$, $E[X] = np$, $Var[X] = np(1-p)$, $\sigma = \sqrt{np(1-p)}$, $\varphi(t) = (q + pe^t)^n$ dove $q = 1-p$. $X_1, \dots, X_n \sim B(1, p)$ indipendenti implica $S = X_1 + \dots + X_n \sim B(n, p)$.
- $X \sim \mathcal{P}(\lambda)$: $P(X = k) = e^{-\lambda} \frac{\lambda^k}{k!}$, $E[X] = \lambda$, $Var[X] = \lambda$, $\sigma = \sqrt{\lambda}$,
 $\varphi(t) = e^{\lambda(e^t - 1)}$. Se $np_n = \lambda$ allora $\lim_{n \rightarrow \infty} \binom{n}{k} p_n^k (1-p_n)^{n-k} = e^{-\lambda} \frac{\lambda^k}{k!}$.

- $X \sim N(\mu, \sigma^2)$: $f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$. $E[X] = \mu$, $Var[X] = \sigma^2$, $\varphi(t) = e^{\mu t} e^{-\frac{t^2\sigma^2}{2}}$. X, Y gaussiane indipendenti, $a, b, c \in \mathbb{R}$ implica $aX + bY + c$ gaussiane. $X \sim N(\mu, \sigma^2)$ si può scrivere come $X = \sigma Z + \mu$, con $Z \sim N(0, 1)$. $F_{\mu, \sigma^2}(x) = \Phi\left(\frac{x-\mu}{\sigma}\right)$. $\Phi(-x) = 1 - \Phi(x)$. $q_\alpha = -q_{1-\alpha}$. Soglie $\mu \pm \sigma q_\alpha$.
- $X \sim Exp(\lambda)$: $f(x) = \lambda e^{-\lambda x}$ per $x \geq 0$, zero per $x < 0$. $E[X] = \frac{1}{\lambda}$, $Var[X] = \frac{1}{\lambda^2}$, $\sigma = \frac{1}{\lambda}$, $\varphi(t) = \frac{\lambda}{\lambda-t}$ per $t < \lambda$. $F(x) = 1 - e^{-\lambda x}$ per $x \geq 0$, zero per $x < 0$.
- TLC: $P\left(\frac{X_1 + \dots + X_n - n\mu}{\sqrt{n}\sigma} \in A\right) \approx P(Z \in A)$, con $Z \sim N(0, 1)$.
- $\bar{X} = \frac{X_1 + \dots + X_n}{n} \sim N\left(\mu, \frac{\sigma^2}{n}\right)$. $E[S^2] = \sigma^2$. $\frac{S^2}{\sigma^2}(n-1) \sim \chi_{n-1}^2$.
- $\mu = \bar{X} \pm \frac{\sigma q_{1-\frac{\alpha}{2}}}{\sqrt{n}}$; $\mu = \bar{X} \pm \frac{S \cdot t_{1-\frac{\alpha}{2}}^{(n-1)}}{\sqrt{n}}$.
- $\left|\frac{\bar{x} - \mu_0}{\sigma} \sqrt{n}\right| > q_{1-\frac{\alpha}{2}}$. $\left|\frac{\bar{x} - \mu_0}{S} \sqrt{n}\right| > t_{1-\frac{\alpha}{2}}^{(n-1)}$. $P\left(|Z| > \left|\frac{\bar{x} - \mu_0}{S} \sqrt{n}\right|\right)$, $P_\mu\left(\bar{X} \in \left[\mu_0 - \frac{\sigma q_{1-\frac{\alpha}{2}}}{\sqrt{n}}, \mu_0 + \frac{\sigma q_{1-\frac{\alpha}{2}}}{\sqrt{n}}\right]\right)$. $\frac{S^2}{\sigma^2}(n-1) > \chi_{\alpha, n-1}^2$. $T = n \sum_{i=1}^k \frac{(\hat{p}_i - p_i)^2}{p_i} = \sum_{i=1}^k \frac{(\hat{X}_i - np_i)^2}{np_i} > \chi_{\alpha, k-1}^2$.
- $y = A + Bx$, $B = \frac{\widehat{Cov}}{S_x^2} = r \frac{S_y}{S_x}$, $\bar{y} = A + B\bar{x}$.